1 2	Duration of obesity exposure between ages 10-40 years and its relationship with cardiometabolic disease risk factors: a cohort study
3	Running title: Obesity duration and cardiometabolic disease risk factors
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23	Word count: 5165
24	Number of tables: 1
25	Number of figures: 4

26 Abstract:

Background: Individuals with obesity do not represent a homogeneous group in terms of
cardio-metabolic risk. Using three nationally representative British birth cohorts, we
investigated whether the duration of obesity was related to heterogeneity in cardiometabolic
risk.

Methods and Findings: We used harmonised body mass index and cardiometabolic 31 32 disease risk factor data from 20 746 participants (49.1% male and 97.2% White British) enrolled in three British birth cohort studies: the 1946 National Survey of Health and 33 Development (NSHD), the 1958 National Child Development Study (NCDS) and the 1970 34 British Cohort Study (BCS70). Within each cohort, individual life course body mass index 35 trajectories were created between 10-40 years of age and from these, age of obesity onset, 36 duration spent obese (range 0-30 years) and cumulative obesity severity were derived. 37 38 Obesity duration was examined in relation to a number of cardiometabolic disease risk 39 factors collected in mid-adulthood: systolic (SBP) and diastolic blood pressure (DBP), high-40 density-lipoprotein cholesterol (HDL-C) and glycated haemoglobin (HbA1c). 41 A greater obesity duration was associated with worse values for all cardiometabolic disease

42 risk factors. The strongest association with obesity duration was for HbA1c: HbA1c levels in those with obesity for <5 years were relatively higher by 5% (95% CI: 4, 6), compared to 43 never obese, increasing to 20% (95% CI: 17, 23) higher in those with obesity for 20-30 44 years. When adjustment was made for obesity severity, the association with obesity duration 45 was largely attenuated for SBP, DBP and HDL-C. For HbA1c however, the association with 46 obesity duration persisted, independent of obesity severity. Due to pooling of three cohorts 47 and thus the availability of only a limited number harmonised variables across cohorts, our 48 49 models included adjustment for only a small number of potential confounding variables, 50 meaning there is a possibility of residual confounding.

51 **Conclusions:** Given that the obesity epidemic is characterised by a much earlier onset of 52 obesity and consequently a greater lifetime exposure, our findings suggest that health policy 53 recommendations aimed at preventing early obesity onset, and therefore reducing lifetime 54 exposure, may help reduce risk of diabetes, independently of obesity severity. However, to 55 test the robustness of our observed associations, triangulation of evidence from different 56 epidemiological approaches (e.g. Mendelian Randomization and negative control studies) 57 should be obtained.

58

59 Author summary:

60 Why Was This Study Done?

- 61 People with obesity (body mass index>30kg/m²) do not all share the same risk for
- 62 development of cardiometabolic disease risk factors.
- 63 The duration a person has spent with obesity over their life course could be one factor
- 64 contributing to the variation observed in cardiometabolic risk.
- 65 However, previous studies have been unable to adequately separate the effects of obesity
- 66 duration (how long a person has been obese) and obesity severity (the magnitude of a
- 67 person's BMI).

68 What Did the Researchers Do and Find?

- 69 We derived body mass index trajectories between 10 and 40 years of age in 20 746
- 70 participants and calculated each person's total time spent with obesity (duration) as well as
- 71 their severity of obesity.
- 72 We related obesity duration to cardiometabolic disease risk factors (systolic and diastolic
- blood pressure, high-density lipoprotein cholesterol and glycated haemoglobin) in mid
- 74 adulthood.
- 75 A greater obesity duration was associated with worse values for all cardiometabolic disease
- risk factors. The strongest association with obesity duration was for HbA1c: HbA1c levels in
- those with obesity for <5 years were relatively higher by 5% (95% CI: 4, 6), compared to
- never obese, increasing to 20% (95% CI: 17, 23) higher in those with obesity for 20-30
- 79 years.
- 80 This positive association between obesity duration and cardiometabolic disease risk factors
- 81 was largely attenuated when adjusting for obesity severity, except for glycated haemoglobin.

82 What Do These Findings Mean?

- The obesity epidemic is characterised by trends towards earlier onset and consequently greater lifetime exposure.
- Our findings are important as they suggest that health policy recommendations aimed at
 preventing early onset obesity, and therefore reducing lifetime obesity exposure, may help
- 87 reduce the risk for diabetes.
- 88 However, due to the small number of potential confounding variables we were able to
- 89 include in our analysis, the contribution of residual confounding to our findings should be
- 90 acknowledged. Furthermore, the robustness of the observed associations should be tested
- 91 using different markers of glucose metabolism and triangulated using different

- 92 epidemiological approaches underpinned by different assumptions and sources of bias (e.g.
- 93 Mendelian Randomization and negative control studies).

94 Introduction:

95 Obesity is a global public health concern. Worldwide prevalence of child and adolescent obesity (defined according to a BMI> 2 standard deviations above age-specific World Health 96 97 Organization cut-offs) has increased from 0.9% and 0.7% in boys and girls, respectively, in 1975 to 7.8% and 5.6%, respectively, in 2016. These increases in child obesity accompany 98 99 significant increases in global adult obesity, with prevalence increasing from 3% and 6.6% of 100 males and females, respectively, in 1975 to 11.6% and 15.7%, respectively, in 2016 (defined 101 according to a body mass index>30kg/m²) [1]. While this epidemic is associated with many adverse health outcomes, particularly cardiovascular disease-related morbidity and mortality 102 [2], individuals with obesity do not represent a homogeneous group in terms of cardio-103 104 metabolic risk. Indeed, there exists a group of individuals who, whilst exceeding the standard BMI cut-off for obesity (\geq 30 kg/m²), are regarded as metabolically healthy because they have 105 an absence of other major cardiovascular risk factors. The life course traits contributing to 106 107 this heterogeneity in cardiometabolic risk have received little attention, but it is likely that a 108 large proportion of the heterogeneity is related, in particular, to the length of time a person 109 spends obese [3,4]. It has been demonstrated that younger individuals are now 110 accumulating greater exposure to overweight or obesity throughout their lives [5], so a comprehensive understanding of the influence of the duration of obesity on the development 111 of cardiometabolic risk factors is critical. 112

Abraham et al [6] published one of the first studies investigating this heterogeneity in 113 cardiometabolic risk for a given weight, observing that rates of some cardiovascular 114 diseases were highest among individuals who were most overweight in adulthood but below 115 116 average weight in childhood. As this study, and others which have replicated that analysis [7–10], are based on weight status at just two time points however, obesity duration can be 117 estimated only crudely. More frequent longitudinal measurements of weight are required for 118 a fuller picture. Furthermore a detailed measurement schedule is also required in order to 119 120 differentiate between obesity duration, the age of obesity onset, and the severity of obesity, as these, though correlated, may confer different health risks [11,12]. For example, due to 121 122 the changes in insulin sensitivity that occur during pubertal development [13], an obesity 123 onset in adolescence may be more deleterious for insulin resistance and diabetes than an 124 onset during another period of the life course.

A handful of studies with such data have observed positive associations between obesity
duration and several cardiometabolic disease risk factors including metabolic syndrome
[11,14], hypertriglyceridemia [14], dyslipidaemia [14,15] and blood pressure [16]. Most
evidence relates to the association with type 2 diabetes however, with numerous studies

129 observing a positive relationship with obesity duration [15,17–23]. The largest of these 130 studies (n=61,821) [21] observed that for each 2-year increment in obesity duration, the risk 131 of type 2 diabetes increased by 14%, though, as observed in other studies [19,22], estimates were attenuated upon adjustment for current weight (representing obesity severity). 132 However, these studies have important limitations, including retrospective designs [14,15], 133 categorising the outcome variable (thus ignoring the observed distribution) [14,15,17,20,21]. 134 an a priori assumption of a linear relationship between obesity duration and outcomes 135 [17,24], and assuming that once a person becomes obese they remain obese, thus 136 137 removing the possibility for weight-cycling [14,15,21–23]. Another important limitation is adjustment of the obesity duration-outcome relationship for current BMI (i.e. at outcome 138 assessment) in order to separate the effects of obesity duration and severity [15,20-22,24]. 139 BMI at outcome assessment does not capture the true extent of obesity severity as it ignores 140 141 (potentially greater) severity occurring at earlier time points. For example, consider two adults, adult A and adult B, who both have a BMI of 35 kg/m² at follow-up and who have 142 both been obese for 20 years. Adult A has had a constant BMI of 35 kg/m², whilst adult B 143 144 has had a BMI as high as 45 kg/m² during this period. It is unlikely, that the cardiometabolic 145 health risks associated with these two profiles are homogeneous.

Recently the concept of 'obese-years' has been proposed, which combines the degree and 146 duration of obesity into a single variable [25,26]. In the study by Araujo et al (2017) [26], an 147 area under the curve of body mass index (BMIAUC) was used to summarize duration and 148 severity of BMI. It is possible however, to obtain a mean obesity severity over any period by 149 150 dividing this AUC by obesity duration, thus separating the effects of severity and duration. To our knowledge, no study has done this and thus robust evidence of the association between 151 152 obesity duration and cardiometabolic risk factors, which is truly independent of obesity severity, is lacking. 153

Using data from three British birth cohort studies, the aim of the present study was to model serial measurements of BMI obtained across the life course in order to derive, for each individual, the following obesity traits: duration of obesity exposure between ages 10-40 years, age of obesity onset and obesity severity. These parameters were then used to relate obesity duration, with and without adjustment for obesity severity, to systolic (SBP) and diastolic blood pressure (DBP), high-density lipoprotein cholesterol (HDL-C) and glycated haemoglobin (HbA1c) in mid-adulthood.

161 **METHODS**

162 Samples

163 The three British birth cohort studies used in these analyses have been previously described

in detail elsewhere [27–29] and were designed to be nationally representative when initiated.

165 The MRC National Survey of Health and Development (NSHD) was initiated in 1946 and

recruited 5 362 participants. The National Child Development Study (NCDS) was initiated in

167 1958 and recruited 17 416 participants. The 1970 British Cohort Study (BCS70) was initiated

- in 1970 and recruited 16 571 participants.
- 169 Ethics statement: All of the studies have received ethical approval and obtained informed
- 170 parental and/or participant consent, both of which cover the secondary analyses reported
- 171 here. Data collection in the NSHD received multicentre research ethics committee approval
- 172 (MREC98/1/121), the NCDS obtained ethical approval from the South East MREC
- 173 (ref:01/1/44) and the BCS70 obtained ethical approval from the National Research Ethics
- 174 Service (NRES) Committee South East Coast Brighton and Sussex (Ref. 15/LO/1446).
- 175 Further details are available from the study websites and/or cohort profiles [27–30].
- 176 For this analysis, we identified a target sample of 20 746 (NSHD: n=2 968; NCDS: n=9 302;
- BCS: n=8 476) participants who attended the biomedical sweep where cardiometabolic
- disease risk factor data were collected (see below) and contributed BMI data for the
- derivation of our exposure variable: obesity duration (S1 Fig).

180 Serial BMI data

- 181 As described elsewhere [31], serial BMI (kg/m²) was derived and harmonised in each study
- 182 from measured or self-reported weight and height collected at the target ages 11, 15, 20
- 183 (self-report), 26 (self-report), 36 and 43 years in the 1946 NSHD; 11, 16, 23 (self-report), 33
- and 42 (self-report) years in the 1958 NCDS; and 10, 16 (one-third self-report), 26 (self-
- report), 30 (self-report), 34 (self-report) and 42 (self-report) years in the 1970 BCS.
- 186 There were 21 009 observations of BMI from 4 702 participants in the 1946 cohort, with 74%
- of the sample having four or more observations. There were 57 545 observations of BMI
- 188 from 16 274 participants in the 1958 cohort, with 80% of the sample having three or more
- observations. Finally, there were 56 275 observations of BMI from 15 437 participants in the
- 190 1970 cohort, with 72% of the sample having three or more observations.

191 Cardiometabolic disease risk factors in adulthood

- 192 In each cohort, a biomedical sweep, with venous blood sampling was conducted in
- adulthood, at 53 years in the 1946 cohort (n=3 053), 44 years in the 1958 cohort (n=9 377)

- and 46 years in the 1970 cohort (n=8 581). Measurements of systolic (SBP) and diastolic
- 195 blood pressure (DBP) were obtained as well as blood cardiometabolic biomarkers (glycated
- 196 haemoglobin (HbA1c) and high-density lipoprotein cholesterol (HDL-C)). More information
- about the measurement protocols can be found in S1 Text.

198 Statistical analysis

TN and WJ determined which analyses to perform and include in the present paper in 199 200 January 2019 after discussing options with all co-authors. The analysis plan was revised in 201 May (modelling obesity duration as a categorical variable rather than a continuous variable) 202 and October 2019 (removing LDL-cholesterol as an outcome due to high amount of missing 203 data) when further exposure and outcome data were obtained and explored. Further analyses were added in June 2020 in response to reviewer comments (adjusting for further 204 putative confounding variables in the regression models, adding sexXduration interaction 205 206 models in the supplementary analyses).

207 Obesity duration parameters

208 In order to identify obesity and derive obesity parameters throughout the life course, we 209 modelled individual child-adulthood trajectories of BMI from 10-40 years of age. These life 210 course BMI trajectories were modelled within each cohort separately, due to the previously 211 described between-cohort heterogeneity in the age-related progression of obesity from 212 childhood to adulthood [5]. Models included all participants who contributed at least one measurement of BMI during the studied age range (NSHD: 11-43 years; NCDS: 11-42 213 years; BCS: 10- 42 years). The BMI trajectories were modelled using restricted cubic splines 214 215 with mixed effects, with measurement occasion at level 1 and individuals at level 2. The 216 restricted cubic splines split the trajectories into piecewise functions of age separated by 217 'knots'. Between the adjacent knots, separate cubic polynomials were fitted, with the spline 218 terms constrained to be linear in the two tails. The number of knots (using the default knot 219 positions as proposed by Harrell[32]) was chosen based on the Bayesian Information 220 Criterion (BIC), with a lower BIC indicating a better fitting model. Once the best fitting model was identified, sex was added as a fixed effect and as interaction terms with the age terms 221 identified in the previous step. Finally, an adjustment for level-1 variation was included to 222 223 allow for differing error associated with measured versus self-reported BMI. From these models, fitted annual-BMI values between 10-40 years were obtained for each individual. 224 Using these fitted BMI values, z-scores were created relative to the International Obesity 225

- Task Force (IOTF) reference [33]. Obesity was defined as a z-score >2.288 in males and
- 227 >2.192 in females, which corresponds to a BMI value of 30 kg/m² at 18 years. Using the sex-
- 220 analis abasity out off asymptotic parameters were derived for each individual. Firstly

229 the presence of obesity at any timepoint was identified, representing any BMI z-score which 230 exceeded the obesity threshold. Secondly, the 'number of times obese' was calculated as 231 the number of times an individual's BMI z-score crossed upwards through the obesity threshold. Thirdly, 'age first obese' was derived, representing the age, in years, when BMI z-232 233 score first crossed upwards through the obesity threshold. 'Total duration of obesity' was 234 calculated as the length of time, in years, that a person's BMI z-score exceeded the obesity threshold; these values were categorised as 0: never obese; 1: obesity 0.01- <5 years; 2: 235 obesity 5 - <10 years; 3: obesity 10 - <15 years; 4: obesity 15 - <20 years; 5: obesity 20+ 236 237 years. Finally, we used the composite trapezoid rule to derive a cumulative obesity severity variable, represented in Fig 1 by the area under the curve and above the obesity threshold. 238 Severity here is expressed in BMI-years above the obesity threshold, reflecting the fact that 239 it incorporates both duration of obesity and the extent to which BMI exceeded the age-240 specific obesity threshold. If this is then divided by obesity duration, it can be interpreted as 241 the 'average obesity severity', i.e. the mean excess BMI above the obesity cut-off. 242

243 Linking obesity parameters to cardiometabolic disease risk factors

244 Pre-specified constants were added to the cardiometabolic disease risk factors to adjust for

being on medication, which has been found to reduce bias [34,35]. The constants were

+10mmHg and +5mmHg for SBP and DBP, respectively, -5% for HDL-C and +1% (absolute)

for HbA1c, obtained from meta-analyses of the effect of blood pressure lowering [36], lipid-

regulating [37–39] and diabetes [40] medications on the respective cardiometabolic riskfactors.

Multiple linear regression was used to relate obesity parameters to the continuous 250 251 cardiometabolic risk factors. As uncertainty in estimated obesity parameters are not taken 252 into account in the confidence intervals for their associations with these continuous 253 cardiometabolic risk factors, standard errors may be underestimated. To correct for this, 254 robust standard errors in these subsequent models were estimated. Data were pooled 255 across cohorts and sexes, thus enabling adjustment of the association between obesity 256 duration and cardiometabolic risk factors for cohort and sex. As HDL-C and HbA1c required 257 transformation to achieve normal distributions, for consistency, we transformed all continuous cardiometabolic risk factors to the 100 loge scale, so that the regression 258 259 coefficients are in units of percentage difference in cardiometabolic risk factor per unit 260 difference in covariate [41]. In a first set of models, the binary variable ever (vs never) obese (between 10-40 years) was tested for association with each cardiometabolic risk factor. In a 261 subsequent set of models, we related the categorical obesity duration variable to each 262 263 cardiometabolic risk factor, with never obese the referent group. The above steps were

- 264 unadjusted for covariates. A subsequent model included adjustments for sex, cohort, birth 265 weight (kg), ethnicity (white vs non-white), social class in childhood (father's social class reported when the child was 10-11 years and according to the Registrar General's Social 266 Classes schema- see S2 Text for more details) and age at follow-up. A final model also 267 268 included an adjustment for average obesity severity. In order to address missingness in 269 covariate data, we used multiple imputation by chained equations (MICE) [42] combining estimates using Rubin's rules [43]. The number of imputations required to achieve 270 convergence of parameter estimates was determined as 100*fraction missing information 271 272 (FMI) [44].
- In addition, in order to aid presentation, we repeated the above steps for a number of
 derived dichotomous cardiometabolic disease risk factor variables, using generalised linear
- models (Poisson distribution with robust error variances) to estimate relative risks (RRs) for
- each outcome. The derived cardiometabolic disease risk factor variables were: hypertension
- 277 (SBP>140mmHg and/or DBP>90mmHg or reported use of BP lowering medication), low
- HDL-cholesterol (<1.03mmol/L in males and <1.29mmol/L in females [45] or reported use of
- 279 lipid-regulating medication) and elevated HbA1c (>5.7% [46] or reported use of diabetes280 medication).
- .
- 281 Beta coefficients from these regression models, i.e. percentage change for continuous
- variables and RRs for binary variables, were plotted. Each figure is split into two, with the
- left-hand side (model 1) showing the estimates from the regression of ever obese (vs never)
- and the right-hand side (model 2) showing estimates of the categorical obesity duration
- 285 variable (vs never).
- 286 Sensitivity analyses

287 First, we repeated the analyses excluding the NSHD cohort as the biomedical sweep 288 occurred much later in this cohort compared to NCDS and BCS70 cohorts, which may have 289 resulted in an underestimation of the association between obesity duration and cardiometabolic disease risk factors. In a related sensitivity analyses, we also replaced the 290 NSHD blood pressure variables to those collected at the 43-year sweep in order to align with 291 the timing of blood pressure measurements in the NCDS and BCS70. No other outcome 292 293 data were available at that age in NSHD however. To identify the extent to which relationships were sex-specific, we also repeated the analyses including a 'sex X obesity 294 duration' interaction. We also performed an analysis which was restricted to those who 295 296 remained obese, assuming that relationships would strengthen when limited to those with 297 persistent obesity and not cycles of obesity.

- Analyses were performed in Stata version 15 (Stata Corp, College Station, TX) and R
- 299 version 3.5.3 (R Core Team 2019).
- 300 This study is reported as per the Strengthening the Reporting of Observational Studies in
- 301 Epidemiology (STROBE) guideline (S1 Checklist).

302 Code availability

- 303 The statistical code for the analyses in this paper has been placed in GitHub, the open-
- 304 access online repository (repository URL: https://github.com/tomnorris1988/Obesity-
- 305 duration-and-cardiometabolic-outcomes).

306 **RESULTS**:

- 307 Descriptive statistics of the cohorts are shown in Table 1. 49.1% of the sample were male
- and 97.2% were White British. As shown in Table 1, the prevalence of 'ever obese' between
- 10-40 years was approximately three times higher in the most recent cohort BCS70 (19.7%,
- n=1673), compared to the oldest cohort NSHD (6.6%, n=197). Average age of first obesity
- onset was less in more recent cohorts, with a median of 30.2 years (inter-quartile range
- 312 (IQR): 25.2, 34.1) in the BCS70 compared to 33.4 years (IQR: 27.6, 37.0) in the NSHD.
- Accordingly, duration of obesity was greater in the most recent cohort BCS70: median 9.7
- 314 years (IQR: 5.9, 14.7), compared to NSHD: 6.2 years (IQR: 2.7, 11.8). The negative
- 315 correlation between age of obesity onset and duration of obesity was almost perfectly
- colinear in the more recent cohort (BCS: -0.99; NCDS: -0.95; NSHD: -0.81), indicating
- almost universal persistence of obesity following its onset in BCS70. Fig 1 provides
- examples of the BMI-z-score trajectories and the derived obesity parameters.
- Average BMI at the biomedical sweep was in the overweight category (>25 kg/m²) in all
- three cohorts but was highest in the BCS70 cohort (27.6 kg/m²; IQR: 24.6, 31.5). For all five
- 321 cardiometabolic risk factors, the mean was highest in the NSHD cohort, reflecting the older
- age at follow-up. This was most notable for SBP, with a mean of 136.0 mmHg (SD: 20.1) in
- the NSHD compared to 126.6 mmHg (16.5) and 124.6 mmHg (15.2) in the NCDS and
- BCS70, respectively. This translated to a much higher prevalence of hypertension in the
- NSHD cohort (68.1%) compared to the NCDS (27.8%) and BCS70 cohorts (23.9%). The
- 326 presence of elevated HbA1c was also considerably higher in the NSHD cohort compared to
- 327 the NCDS and BCS70 (35.8% vs 15.0% and 16.5%, respectively).

328 Relationship of obesity parameters with cardiometabolic disease risk factors

- Results from the unadjusted analysis are included in S1 and S2 Tables. Here we report
- estimates from the adjusted analyses, presented in Fig 2-4 and with corresponding
- 331 estimates in S3-S6 Tables.
- 332 HbA1c
- Being ever obese at any age between 10-40 years (versus never obese) was associated
- with an 9.0% (95% CI: 8.2, 9.9) higher HbA1c (Fig 2, left panel), which reduced to 4.5%
- higher (95% CI: 3.5, 5.6) when adjusted for obesity severity. HbA1c increased linearly with
- obesity duration, from 5% excess for <5 years duration up to 19.9% (95% CI: 16.5, 23.3) for
- 20-30 years duration (p(trend) < 0.001). Upon adjustment for obesity severity, the trend
- remained (p(trend)=0.007) but was attenuated, particularly for 20-30 years, which reduced
- from 19.9% to 11.6% (95% CI: 5.9, 17.2), a relative reduction of 42%.

- 340 There was also a linear trend between obesity duration and risk for elevated HbA1c, with
- those obese <5 years having a 2.1 (95% CI: 1.8, 2.4) times higher risk of elevated HbA1c of
- 342 compared to never obese, which more than doubled in those obese for 20-30 years (relative
- risk 4.6; 95% CI: 3.9, 5.5, *p(trend)*<0.001) (Fig 2, right panel). However, upon adjustment for
- obesity severity, this graded relationship was attenuated (p(trend)=0.006).

345 SBP and DBP

- 346 There was a positive relationship between ever being obese between 10-40 years and both
- 347 systolic and diastolic blood pressure. For example, ever obese was associated with a 6.1%
- 348 (95% CI: 5.6, 6.6) higher SBP and 7.1% (95% CI: 6.6, 7.7) higher in DBP at follow-up (vs
- never obese) (Fig 3, panel 1 and 2). Obesity duration was also positively associated with
- both SBP and DBP, such that SBP was 5.0% higher in those who were obese <5 years
- compared to those never obese, increasing to 9.0% higher for 20-30 years (p(trend) < 0.001).
- However, upon adjustment for obesity severity, evidence for this dose-response association
- 353 was greatly reduced (SBP: p(trend)=0.975; DBP: p(trend)=0.294).
- 354 Consistent with these findings, ever being obese between 10-40 years (vs never) was
- associated with a relative risk for hypertension of 1.6 (95% CI: 1.5, 1.7), independent of
- obesity severity (Fig 4, panel 1 and S6 Table). For obesity duration, a similar pattern was
- observed to that seen for SBP and DBP, i.e. a gradually increasing risk for hypertension with
- increasing time spent obese (*p(trend*)<0.001), evidence for which weakened when adjusted
- for obesity severity (p(trend)=0.456).
- 360 HDL-cholesterol
- 361 A negative relationship was observed between obesity and HDL-cholesterol, such that
- obesity at any point between 10-40 years was associated with a 16.4% (95% CI: 17.6, 15.2)
- lower HDL-C at follow-up (Fig 3, panel 3), attenuating to 12.3% lower when adjusted for
- 364 severity. There was a linear trend in the effect of obesity duration on HDL-C, such that HDL-
- C levels in those with obesity <5 years were 12.4% (95% CI: 10.4, 14.4) lower than those
- never obese, which increased to 24.8% (95% CI: 20.5, 29.1) lower in those who had been
- obese for 20-30 years (*p(trend)*<0.001). Upon adjustment for obesity severity, evidence for
- the trend attenuated (p(trend)=0.117).
- 369 This resulted in a relative risk for low-HDL-C of 2.0 (95% CI: 1.8, 2.2) in those who were ever
- obese between 10-40 years (vs never), independent of obesity severity (Fig 4, panel 2 and
- S6 Table). For obesity duration there was a linear trend of increasing risk (p(trend) < 0.001),
- 372 which remained on adjustment for severity, though evidence for this was attenuated
- 373 (*p(trend)*=0.037).

- 374 Sensitivity analysis
- 375 Similar results were found when the analysis was limited to the NCDS and BCS70 cohorts
- 376 (S7 and S8 Tables), thus accounting for the difference in the age at follow-up in the NSHD.
- 377 Similarly, replacing the blood pressure variables in the NSHD cohort with those collected at
- the age 43-year sweep, in order to be more aligned with the age at follow-up in the NCDS
- and BCS70, did not change results (S9 and S10 Tables). When stratified by sex,
- associations were consistently stronger in females (S11 and S12 Tables) and especially for
- the dichotomous cardiometabolic disease risk factor variables. Finally, estimates were
- 382 largely unchanged when the analysis was limited to those with persistent obesity (i.e. staying
- 383 obese after first onset) (S13 and S14 Tables).

384 **DISCUSSION**:

This study utilised longitudinal BMI data from three British birth cohort studies to model each 385 person's obesity history and derive individual obesity parameters. Ever being obese between 386 10-40 years of age, compared to never being obese, was associated with less favourable 387 388 levels of all cardiometabolic disease risk factors. More time spent obese was associated with worse profiles for all cardiometabolic disease risk factors, though greatest for HbA1c. When 389 390 adjustment was made for obesity severity, the strength of the evidence in support of an 391 association between obesity duration and SBP, DBP and HDL-C was weak (p>0.1). For HbA1c however, though the association with obesity duration also attenuated when 392 adjusting for obesity severity, the strength of evidence remained strong. The study design, in 393 394 particular the fact that most individuals who became obese remained obese, has meant that age of obesity onset and obesity duration are very highly negatively correlated. Our results 395 also therefore mean that, after accounting for obesity severity, an earlier age of obesity 396 397 onset was only associated with HbA1C. These key findings were robust to a range of 398 sensitivity analyses.

399 In attempting to separate the effects of obesity duration and severity on cardiometabolic 400 health, previous studies have simply adjusted for BMI (or waist circumference) at the time of outcome assessment [14,15,19,21,22,47]. This, however, only provides an indication of 401 402 obesity severity at that particular point in time. Our study represents an advance over these 403 studies however, as we have been able to measure obesity severity accumulated over the life course, and by adjusting this for the time spent obese we have been able to 404 appropriately separate the effects of obesity duration and severity. As such, these findings 405 406 provide novel, robust evidence regarding the independent association of obesity duration with cardiometabolic disease risk factors. 407

408 Our findings are in line with other studies which have observed an attenuated, but persisting, effect of obesity duration on diabetes risk or impaired glucose metabolism, once obesity 409 severity is accounted for [15,19,21,22]. In another NCDS analysis (n=7855), Power et al 410 411 (2011) [19] observed that compared to those never obese, those with the greatest duration 412 of obesity (i.e. onset <16 years), had an almost 24-fold increased risk of having HbA1c >7% (and/or a diagnosis of diabetes) at 45 years. While this risk was substantially attenuated 413 414 upon adjustment for current BMI, it still remained over 4 times higher compared to those 415 never obese. In addition we have observed, in line with Pontirolli et al (1998) [15], a specific effect of obesity duration on glucose metabolism. In their study of 760 obese adults (average 416 age 51 years) obesity duration was a risk factor for glucose intolerance and type 2 diabetes 417 418 but not for hypertension or hyperlipidaemia [15]. Evidence in support of our finding of no

independent association of obesity duration with HDL-C is lacking. To our knowledge only
one other study has investigated this and observed an association in females only, though
the strength of evidence was modest (p=0.05) [14].

422 In addition to the cited empirical studies, there is also a plausible biological mechanism 423 supporting the observed association between obesity duration and HbA1c (reflecting 424 impaired glucose metabolism). Obesity is characterised by enlarged fat stores, which results 425 in enhanced lipolysis and an increase in circulating free fatty acids. This state leads to 426 peripheral and hepatic insulin resistance [48,49], resulting in a compensatory insulin hypersecretion by the pancreatic β -cells in order to preserve normoglycemia[50]. Prolonged 427 obesity leads to β-cell exhaustion [51], culminating in a reduced insulin response and an 428 429 inability to maintain normoglycemia [52]. In addition, prolonged obesity may represent a state in which subcutaneous adipose stores have been exhausted, with the consequence 430 being a deposition of adipose tissue around the visceral organs (e.g. liver and pancreas), 431 432 with fat stored in these areas (i.e. 'ectopic fat') being strongly related to insulin resistance

433 [53].

Despite the persisting independent effect of obesity duration on HbA1c levels, a substantial 434 435 reduction in the effect was observed once severity of obesity had been accounted for. This suggests that in those who have been exposed to obesity for a prolonged period, there is still 436 437 opportunity to return to more favourable HbA1c levels if a degree of weight loss is achieved. 438 For example, upon adjustment for severity, the risk of elevated HbA1c in those who had been obese for 20-30 years reduced from more than a 4-fold increased risk (relative to never 439 obese), to a level similar to those obese for half as long, i.e. 10-15 years (RR: 3.0; 95%CI: 440 441 2.3, 4.0).

442 There was some evidence that the association between obesity duration and the

dichotomous cardiometabolic outcomes was stronger in females than males (S11 and S12

Tables). Sex-specific associations have been observed in other studies [14,22,54]. Janssen

et al (2004) [14] for example, observed an independent effect of overweight/obesity duration

446 on risk for insulin resistance and type 2 diabetes (and also hypertension,

447 hypertriglyceridemia, low-HDL-C and metabolic syndrome) in females, but not in males

448 except for hypertriglyceridemia. A sex difference was also observed in the Framingham

Heart Study [54]. Tanamas et al (2015) observed an association between obesity duration

and risk for hypertension in females but not in males (ages 30-62 years) [54]. As

451 summarised in the review by Jarvis (2015) [55], there are fundamental differences in the

452 control of metabolic homeostasis between males and females. Females are more likely to

453 gain fat, and though abdominal obesity more commonly affects males than females, the

- 454 prevalence of abdominal obesity has increased more in females than in males [56].
- 455 Furthermore, the prevalence of visceral obesity associated with metabolic syndrome is two
- to ten times higher in women throughout the world [57–59]. It may be therefore, that
- 457 compared to males, females are more exposed to this metabolically-volatile adipose tissue
- 458 and thus at increased risk of its deleterious outcomes.

459 Strengths

The key strength of our study is the derivation, using over 130 000 serial BMI observations across the life course, of individualised obesity parameters which enabled us to distinguish between obesity severity and duration. In addition, the pooling of data from three nationally representative cohorts means the observed associations are based on a far larger sample than most previous studies and are likely to be generalisable to the underlying population.

465 Limitations

Our definition of obesity was based on BMI, which despite exhibiting a strong positive 466 correlation with direct estimates of fat mass [60], is only an indicator of total body adiposity. 467 468 However, it remains the most commonly used, widely accepted, and practical measure of obesity in both children and adults. Our trajectories were dependent on the frequency of BMI 469 470 measurements across the life course, with some intervals spanning 10 years. As such we 471 may not have captured instances of weight cycling between measurement occasions. 472 Measurement protocols for weight and height were not consistent within and between studies, which may have introduced bias if, self-reported measurements were systemically 473 474 under or over-reported. It has been shown that people with greater BMIs tend to under-475 report their weight [61,62], suggesting that estimates of obesity duration (and severity) may 476 be conservative in our study. Our regression models included adjustment for only a small 477 number of covariates, which means there is a possibility of residual confounding. As we 478 have combined three cohorts, any included variable must be harmonised across each cohort 479 so that the variable conveys the same thing in each cohort. This is only the case for a small 480 number of variables in the cohorts we have used. As all of the included studies suffered from attrition, which is more extensive in those from lower SEP groups and/or with higher BMI 481 [63,64], we may have inadvertently selected a more socioeconomically advantaged and 482 483 thinner sample which in addition to a loss of power, may have introduced bias into the observed associations. In addition, as the NSHD, NCDS and BCS70 cohorts are either 484 exclusively (NSHD), or predominantly White British, we are unable to generalise the results 485 486 to other ethnic groups. Finally, the biomedical sweep in the NSHD cohort was conducted 9 and 7 years later than the NCDS and BCS cohorts, respectively, which may impair cross-487 cohort comparability (underpinning the decision to pool cohorts). However, supplementary 488

analyses limited to the NCDS and BCS cohorts only and replacing the NSHD blood pressure
variables with those collected at 43 years, produced similar estimates (S7-S10 Tables).

Associations observed in this study suggest that there are benefits in delaying the onset of 491 492 obesity, as risks of elevated HbA1c were positively associated with time spent with obesity, 493 independent of the degree of severity. Interventions aiming to prevent childhood obesity 494 therefore have the potential to reduce the long-term risk of developing diabetes. However, 495 we also observed an amelioration of HbA1c profiles in those who had been exposed to 496 obesity for a prolonged period, once severity of obesity is accounted for. As such, people with obesity should be encouraged to lose weight in order to return their HbA1c levels to 497 more favourable values. Firstly however, more research using different epidemiological 498 499 approaches underpinned by different assumptions and sources of bias (e.g. Mendelian Randomization and negative control studies) is needed to test the robustness of these 500 501 findings.

502 Conclusion

We found a dose response relationship between the duration of obesity and HbA1c. 503 504 independent of obesity severity. Given that the obesity epidemic is characterised by trends 505 towards earlier onset and consequently greater lifetime exposure, our findings are important 506 as they suggest that health policy recommendations aimed at preventing early onset obesity, 507 and therefore reducing lifetime obesity exposure, may help reduce the risk for diabetes. For those who are already obese, reducing obesity severity can also improve their metabolic 508 profile. Accordingly, prevention strategies could consider both the duration and severity of 509 510 obesity.

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Supplemental files

S1 Checklist: STROBE checklist

S1 Text: Measurement protocol for collection of cardiometabolic outcomes in adulthood

S2 Text: Childhood social class

S1 Table: Association between ever obese and categories of obesity duration (vs never obese) and cardiometabolic disease risk factors*† (imputed, unadjusted)

S2 Table: Association between ever obese and categories of obesity duration (vs never obese) and dichotomous cardiometabolic outcomes (imputed, unadjusted)

S3 Table: Association between ever obese and categories of obesity duration (vs never obese) and cardiometabolic disease risk factors*† (imputed, adjusted for sex, cohort, age at follow-up, ethnicity, birth weight and childhood social class)

S4 Table: Association between ever obese and categories of obesity duration (vs never obese) and dichotomous cardiometabolic outcomes (imputed, adjusted for sex, cohort,age at follow-up, ethnicity, birth weight and childhood social class)

S5 Table: Association between ever obese and categories of obesity duration (vs never obese) and cardiometabolic disease risk factors*† (imputed, adjusted for sex, cohort, age at follow-up, ethnicity, birth weight, childhood social class and obesity severity)

S6 Table: Association between ever obese and categories of obesity duration (vs never obese) and dichotomous cardiometabolic outcomes (imputed, adjusted for sex, cohort, age at follow-up, ethnicity, birth weight, childhood social class and obesity severity)

S7 Table: Association between ever obese and categories of obesity duration (vs never obese) and cardiometabolic disease risk factors (imputed, adjusted for sex, cohort, age at follow-up, ethnicity, birth weight, childhood social class and obesity severity): excluding NSHD

S8 Table: Association between ever obese and categories of obesity duration (vs never obese) and dichotomous cardiometabolic outcomes (imputed, adjusted for sex, cohort, age at follow-up, ethnicity, birth weight, childhood social class and obesity severity): excluding NSHD

S9 Table: Association between ever obese and categories of obesity duration (vs never obese) and cardiometabolic disease risk factors (imputed, adjusted for sex,

cohort, age at follow-up, ethnicity, birth weight, childhood social class and obesity severity): using blood pressure at 43 years in NSHD

S10 Table: Association between ever obese and categories of obesity duration (vs never obese) and categorical cardiometabolic outcomes (imputed, adjusted for sex, cohort, age at follow-up, ethnicity, birth weight, childhood social class and obesity severity): using blood pressure at 43 years in NSHD

S11 Table: Association between ever obese and categories of obesity duration (vs never obese) and cardiometabolic disease risk factors (imputed, adjusted for cohort, age at follow-up, ethnicity, birth weight, childhood social class and obesity severity): sex interaction

S12 Table: Association between ever obese and categories of obesity duration (vs never obese) and dichotomous cardiometabolic outcomes (imputed, adjusted for cohort, age at follow-up, ethnicity, birth weight, childhood social class and obesity severity): sex interaction

S13 Table: Association between ever obese and categories of obesity duration (vs never obese) and cardiometabolic disease risk factors (imputed, adjusted for sex, cohort, age at follow-up, ethnicity, birth weight, childhood social class and obesity severity): limited to those who once obese were always obese

S14 Table: Association between ever obese and categories of obesity duration (vs never obese) and dichotomous cardiometabolic outcomes (imputed, adjusted for sex, cohort, age at follow-up, ethnicity, birth weight, childhood social class and obesity severity): limited to those who once obese were always obese

S1 Fig: Sample flow diagram

		NSHD 1946 (n=2968)	1958 NCDS (n=9302)	1970 BCS (n=8476)
Sex				
Males	n (%)	1459 (49.2)	4630 (49.8)	4106 (48.4)
Females	n (%)	1509 (50.8)	4672 (50.2)	4370 (51.6)
Ethnicity				
White British	n (%)	2968 (100)	9089 (97.7)	7882 (93.0)
Other ^a	n (%)	0 (0)	205 (2.2)	376 (4.4)
Missing	n (%)	0 (0)	8 (0.1)	218 (2.6)
Obesity traits				
Never obese	n (%)	2771 (93.4)	8267 (88.9)	6803 (80.3)
Ever obese	n (%)	197 (6.6)	1035 (11.1)	1673 (19.7)
Age first onset (years)	Median (IQR)	33.4 (27.6; 37.0)	31.5 (25.4; 36.1)	30.2 (25.2; 34.1)
Total duration (years)	Median (IQR)	6.2 (2.7; 11.8)	8.3 (3.9; 14.4)	9.7 (5.9; 14.7)
Correlation		-0.81	-0.95	-0.99
(age onset x duration obese)				0.00
Number of periods				
1	n (%)	192 (97.5)	1023 (98.8)	1671 (99.9)
2	n (%)	4 (2.0)	12 (1.2)	2 (0.1)
3	n (%)	1 (0.5)	0	0
Obesity severity (BMI-years)	Median (IQR)	5.6 (1.1; 22.7)	9.4 (1.7; 35.0)	17.1 (4.3; 48.0)
Correlation (duration x severity)		0.86	0.85	0.80
Biomedical outcomes		% missing	% missing	% missing

Table 1 Descriptive statistics for life course obesity parameters and cardiometabolic disease risk factors at the biomedical sweep of those in target study sample (n=20 746)

mean (SD)	-	53.5 (0.2)	-	45.2 (0.4)	-	47.3 (0.7)
median (IQR)	1.3	26.6 (24.2;	1.3	26.6 (24.0; 29.9)	13.4	27.6 (24.6;
		29.9)				31.5)
p(9/)	1.3	707 (24.1)	1.3	2239 (24.4)	13.4	2424 (33.0)
II (%)						
mean (SD)	1.9	136.0 (20.1)	0.5	126.5 (16.5)	11.5	124 6 (15 2)
						124.0 (15.2)
mean (SD)	1.9	84.4 (12.2)	0.5	78.8 (10.8)	11.5	77.3 (11.0)
median (IQR)	20.2	1.6 (1.3; 2.0)	16.1	1.5 (1.3; 1.8)	29.5	1.5 (1.2; 1.8)
n (%)	19.3	312 (13.0)	14.5	1595 (20.1)	28.5	1385 (22.8)
median (IQR)	13.6	5.7 (5.3; 5.9)	15.2	5.3 (5.0; 5.4)	29.9	5.4 (5.3; 5.6)
n (%)	13.2	921 (35.8)	13.8	1206 (15.0)	39.3	987 (16.5)
	mean (SD) median (IQR) n (%) mean (SD) mean (SD) n (%) median (IQR) n (%) median (IQR) n (%)	mean (SD) - median (IQR) 1.3 n (%) 1.3 mean (SD) 1.9 mean (SD) 1.9 mean (SD) 1.9 median (IQR) 20.2 n (%) 19.3 median (IQR) 13.6 n (%) 13.2	$\begin{array}{c cccc} \mbox{mean} ({\rm SD}) & - & 53.5 (0.2) \\ \mbox{median} ({\rm IQR}) & 1.3 & 26.6 (24.2; \\ 29.9) \\ \mbox{n} (\%) & 1.3 & 707 (24.1) \\ \mbox{mean} ({\rm SD}) & 1.9 & 136.0 (20.1) \\ \mbox{mean} ({\rm SD}) & 1.9 & 84.4 (12.2) \\ \mbox{n} (\%) & 1.9 & 1993 (68.1) \\ \mbox{median} ({\rm IQR}) & 20.2 & 1.6 (1.3; 2.0) \\ \mbox{n} (\%) & 19.3 & 312 (13.0) \\ \mbox{median} ({\rm IQR}) & 13.6 & 5.7 (5.3; 5.9) \\ \mbox{n} (\%) & 13.2 & 921 (35.8) \\ \end{array}$	$\begin{array}{c ccccccccccccccccccccccccccccccccccc$	$\begin{array}{c ccccccccccccccccccccccccccccccccccc$	$\begin{array}{c ccccccccccccccccccccccccccccccccccc$

*Original values (i.e. not adjusted for medication use); aOther ethnicities: White other, Mixed race, Indian, Pakistani, Bangladeshi, Other Asian, Caribbean, African, Other Black, Chinese. bHypertension: SBP/DBP≥140/90mmHg and/or on BP lowering medication; Cow-HDL-C: according to NCEP ATPIII criteria and/or on lipid-regulating medication; dElevated HbA1c: according to CDC criteria and/or on diabetes medication

Figure legends:

Fig 1: Example obesity traits (onset, duration and severity (area-under-the curve and above obesity cut-off)) derived from the BMI-z-score trajectories of two random participants



BMI< obesity cut-off BMI> obesity cut-off

Fig 2: Association between ever obese and categories of obesity duration (vs never obese) and HbA1c (left panel) and risk for elevated HbA1c (right panel)



Fig 3: Association between ever obese and categories of obesity duration (vs never obese) and SBP, DBP and HDL-C



Fig 4: Association between ever obese and categories of obesity duration (vs never obese) and risk for hypertension and low HDL-cholesterol

